

# The large-maturity smile for the Stein-Stein model

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**Abstract:** We compute the large-maturity smile for the correlated Stein-Stein stochastic volatility model  $dS_t = S_t Y_t dW_t^1$ ,  $dY_t = \kappa(\theta - Y_t)dt + \sigma dW_t^2$ ,  $dW_t^1 dW_t^2 = \rho dt$ , using the known closed-form solution for the characteristic function of the log stock price given in Schöbel&Zhu[SZ99]. The Stein-Stein model is not covered by the results in [FK13] and [JKRM13] because the volatility fails to satisfy the sublinear growth condition in [FK13] and is not an affine model.†

## 1. Introduction

The last few years have witnessed a number of articles on large-time asymptotics for stochastic volatility models with/without a jump component. Using the Gärtner-Ellis theorem, [FJ11] compute the implied volatility smile for the popular Heston stochastic volatility model when  $\kappa > 0$ ,  $\kappa > \rho\sigma$ , in the large-time, large log-moneyness regime and [FJM10] compute the correction term using saddlepoint methods; the large-time smile is identical to the large-time smile for the Barndorff-Nielsen Normal Inverse Gaussian model, and [GJ11] show that the asymptotic smile can be computed in closed-form via the Gatheral SVI parameterization. [JM12] derive similar results for a displaced Heston model, and relax the aforementioned conditions on  $\kappa, \rho, \sigma$ . [JKRM13] have extended the results in [FJ11] to a general class of affine stochastic volatility models (with jumps), which includes the Heston model with state-independent jumps, the Bates model with state-dependent jumps and the Barndorff-Nielsen-Shephard model.

[FP12] compute large-time asymptotics for the SABR model with  $\beta = 1, \rho \leq 0$  and  $\beta < 1, \rho = 0$ ; in particular for  $\beta = 1, \rho \leq 0$ , they compute a closed-form expression for the asymptotic log stock price density and establish large-time asymptotics for the CEV model and the uncorrelated CEV-Heston model in the large-time, fixed-strike regime and a new large-time, large log-moneyness regime. [Forde11b] derives similar results for the modified SABR model in terms of the large-time asymptotic density of the Brownian exponential functional.

The long-term asymptotic behavior of the smile for exponential Lévy models and more general martingale models have been studied in [RT10], where it is proved that for fixed log-moneyness  $k$  and large maturity, the implied volatility converges to a constant value that does not depend on  $k$ . This phenomenon is typically referred to as the “smile-flattening” effect, which arises from the large deviation principle for i.i.d. random variables (see e.g. Cramér’s theorem in [DZ98]). For a general exponential Lévy model with mild conditions on the cumulant generating function, [GL11] derive an expansion of the form  $\hat{\sigma}_t(x)^2 = \sigma_\infty^2 + a_1(x)/t + a_2(x)/t^2 + o((\log t)^2/t^3)$  as  $t \rightarrow \infty$  for the implied volatility  $\hat{\sigma}_t(x)$  at log-moneyness  $x$  and maturity  $t$ , where  $a_1(x)$  and  $a_2(x)$  are respectively affine and quadratic in  $x$ .

In [Forde11], the author derives a large deviation principle for the log stock price under an uncorrelated stochastic volatility model driven by an Ornstein-Uhlenbeck process with a bounded volatility function. For this we use the fact that the occupation measure for the Ornstein-Uhlenbeck process satisfies an LDP with a good, convex lower semicontinuous rate function under the topology of weak convergence (and under the Prohorov metric), see section 7 in Donsker&Varadhan[DV76] (see also page 178 in Stroock[Str84] and [Pin85]), combined with the standard contraction principle and exponential tightness. In [FK13], we relax the assumptions of bounded volatility and zero correlation made in [Forde11]. The rate function for  $X_t/t$  now has the variational representation  $I(x) = \inf_{\mu \in \mathcal{P}(\mathbb{R})} \frac{(x - M(\mu))^2}{2\nu(\mu)} + I_\alpha(\mu)$ , for some linear functionals  $M, \nu$  which depend on the correlation  $\rho$ . Using the LDP, we translate these results into large-time asymptotics for call options and implied volatility, and we extend the analysis to incorporate stochastic interest rates, by deriving a similar LDP for a three-factor model driven by a CIR short rate process.

In this article, we look at the large-time behavior of the closed form expression for the characteristic function of the log stock price under the Stein-Stein model introduced in [SS91], which is derived in [SZ99]. Using the Gärtner-Ellis

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Theorem from large deviations theory, we compute a large-time large deviation principle for the log stock price. From this we can then characterize the large-time behavior of call option prices and implied volatility in the large-time, large log-moneyness regime. The Stein-Stein model reduces to a special case of the Heston model when the mean reversion level  $\theta = 0$ . We refer the reader to Deuschel et al.[DFJV14] for a discussion on tail asymptotics for the Stein-Stein model using Laplace's method on Wiener space for a small-noise diffusion process and some simple scaling properties, and the earlier work on tail asymptotics for the zero correlation case in Gulisashvili&Stein[GS10].

## 2. Large deviation theory and the Gärtner-Ellis theorem

In this section, we recall some fundamental notions in large deviations theory (we refer the reader to Section 2.3 in [DZ98] and Section 2.2 in [JM12] for more details). A family of random variables  $(Z_t)$  is said to satisfy the large deviation principle (LDP) as  $t \rightarrow \infty$  with good rate function  $I$  if for all  $B \in \mathcal{B}(\mathbb{R})$  we have the following bounds

$$-\inf_{x \in B^o} I(x) \leq \liminf_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(Z_t \in B) \leq \limsup_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(Z_t \in B) \leq -\inf_{x \in \bar{B}} I(x),$$

where  $B^o$  ( $\bar{B}$ ) denotes the interior (resp. closure) of  $B$ .

We now assume that the cumulant generating function  $V_t(p) = \log \mathbb{E}(e^{pZ_t})$  is finite on some neighbourhood of zero and that the following limit exists as an extended real number

$$V(p) = \lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{pZ_t}) \quad \forall p \in \mathbb{R}. \quad (2.1)$$

Let  $\mathcal{D}_V = \{p \in \mathbb{R} : |V(p)| < \infty\}$  and assume that  $\{0\} \in \mathcal{D}_V^o$ . From Hölder's inequality we can show that  $V_t$  is convex for all  $t > 0$  and the limit  $V$  is also convex (see Lemma 2.3.9 in [DZ98]). Moreover  $V(0) = 0$ , thus (by convexity) we see that  $V(p) > -\infty$  for all  $p \in \mathbb{R}$ .  $V : \mathbb{R} \rightarrow (-\infty, \infty]$  is called *essentially smooth* if  $V$  is differentiable in  $\mathcal{D}_V^o$  and satisfies  $\lim_{n \rightarrow \infty} |V'(p_n)| = \infty$  for every sequence  $(p_n)$  in  $\mathcal{D}_V^o$  which converges to a boundary point of  $\mathcal{D}_V^o$ . A cgf  $V$  which satisfies this second property is called *steep*. The Fenchel-Legendre transform  $V^*$  of  $V$  is defined by the variational formula

$$V^*(x) = \sup_{p \in \mathbb{R}} [px - V(p)]$$

for all  $x \in \mathbb{R}$ , with an effective domain  $\mathcal{D}_{V^*} = \{x \in \mathbb{R} : V^*(x) < \infty\}$ . In general  $V^*$  can be discontinuous and  $\mathcal{D}_{V^*}$  can be a strict subset of  $\mathbb{R}$  (see section 2.3 in [DZ98] for some simple examples).

We now state a simplified version of Gärtner-Ellis theorem (c.f. Theorem 2.3.6 in [DZ98]) which will be needed in the next section.

**Theorem 2.1.** *Let  $(Z_t)_{t>0}$  be a family of random variables for which  $V$  as defined in (2.1) satisfies  $\{0\} \in \mathcal{D}_V^o$ . If  $V$  is essentially smooth and lower semicontinuous, then the LDP holds with good rate function  $V^*$ .*

## 3. The Stein-Stein model

From here on, we work on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  with a filtration  $(\mathcal{F}_t)_{t \geq 0}$  throughout, supporting two independent Brownian motions and satisfying the usual conditions. We now recall the Stein-Stein stochastic volatility model for a log stock or forward price process  $X_t = \log S_t$ :

$$\begin{cases} dX_t = -\frac{1}{2}Y_t^2 dt + Y_t dW_t^1, \\ dY_t = \kappa(\theta - Y_t)dt + \sigma dW_t^2 \end{cases} \quad (3.1)$$

where  $\kappa, \sigma > 0$ ,  $X_0 = x_0$ ,  $Y_0 = y_0$ , and  $W^1, W^2$  are Brownian motions such that  $dW_t^1 dW_t^2 = \rho dt$ ,  $|\rho| < 1$ . The law of  $X_t - x_0$  does not depend on  $x_0$ , so without loss of generality we set  $X_0 = 0$ .

We first verify the martingale property for  $S_t$ .

**Proposition 3.1.**  $(S_t)$  is a martingale.

*Proof.* Let  $0 < t_1 < t_2 < \infty$ . We know that  $\sup_{t \geq 0} \mathbb{E}(e^{cY_t^2}) < \infty$  if  $c < \kappa/\sigma^2$ , using that  $Y_t \sim N(e^{-\kappa t}y_0 + \theta(1 - e^{-\kappa t}), \frac{\sigma^2}{2\kappa}(1 - e^{-2\kappa t}))$ . Now consider a uniform random variable  $U$  on  $[t_1, t_2]$ , independent of  $S$ , and let  $\mathcal{F}_t^Y = \sigma(Y_s; 0 \leq s \leq t)$  denote the filtration generated by the  $Y$  process. Then we have

$$\begin{aligned} \mathbb{E}(e^{\frac{1}{2} \int_{t_1}^{t_2} Y_s^2 ds}) &= \mathbb{E}(e^{\frac{1}{2}(t_2-t_1)\mathbb{E}(Y_U^2 | \mathcal{F}_t^Y)}) \\ &\leq \mathbb{E}(\mathbb{E}(e^{\frac{1}{2}(t_2-t_1)Y_U^2} | \mathcal{F}_t^Y)) \\ &\quad \text{(using the conditional Jensen's inequality)} \\ &= \mathbb{E}\left(\frac{1}{t_2-t_1} \int_{t_1}^{t_2} e^{\frac{1}{2}(t_2-t_1)Y_s^2} ds\right) < \infty \\ &= \frac{1}{t_2-t_1} \int_{t_1}^{t_2} e^{\frac{1}{2}(t_2-t_1)\mathbb{E}(Y_s^2)} ds < \infty \\ &\quad \text{(by Fubini's theorem)} \end{aligned}$$

for  $\frac{1}{2}(t_2-t_1) \leq \kappa/\sigma^2$ . By Corollary 5.14, p.199 in [KS91], we conclude that  $S_t = e^{-\frac{1}{2} \int_0^t Y_s^2 ds + \int_0^t Y_s dW_s^1}$  is a martingale.  $\square$

### 3.1. The large-time large deviation principle for the re-scaled log return

The following proposition establishes a large-time large deviation principle for the re-scaled log return for the Stein-Stein model:

**Proposition 3.2.**  $X_t/t$  satisfies a large-time LDP as  $t \rightarrow \infty$  with a good convex continuous rate function given by the Fenchel-Legendre transform

$$I(x) = \sup_p [px - V(p)]$$

where

$$V(p) = V(p; \kappa, \theta, \sigma, \rho) = \begin{cases} \frac{1}{2} [\kappa - p\rho\sigma + \frac{(p-1)p\theta^2\kappa^2}{\Gamma(p)^2} - \Gamma(p)] & (p \in (p_-, p_+)) \\ +\infty & (p \notin (p_-, p_+)) \end{cases}$$

$\bar{\rho} = \sqrt{1 - \rho^2}$ ,  $\Gamma(p) = \sqrt{\kappa^2 - 2p\kappa\rho\sigma + p(1 - p\bar{\rho}^2)\sigma^2}$  and  $p_{\pm} = \frac{\sigma^2 - 2\kappa\rho\sigma \pm \sigma\sqrt{4\kappa^2 - 4\kappa\rho\sigma + \sigma^2}}{2\sigma^2\bar{\rho}^2}$  are the roots of  $\Gamma(p)^2$ .  $I$  is continuous and attains its minimum value uniquely at  $x^* = V'(0) = -\frac{1}{2}(\theta^2 + \sigma^2/2\kappa)$ .

*Proof.* From Eq 13 in [SZ99], we have the following closed-form expression for the characteristic function of the log return

$$\phi_t(u) = \mathbb{E}(e^{iuX_t}) = e^{-\frac{1}{2}iu\rho(\sigma^{-1}y^2 + \sigma t) + \frac{1}{2}D(t, \hat{s}_1, \hat{s}_2, \hat{s}_3)y^2 + B(t, \hat{s}_1, \hat{s}_2, \hat{s}_3)y + C(t, \hat{s}_1, \hat{s}_2, \hat{s}_3)} \quad (3.2)$$

for  $u \in \mathbb{R}$ , where

$$\begin{aligned} D(t, T) &= \frac{1}{\sigma^2} \left[ \kappa - \gamma_1 \frac{\sinh(\gamma_1 t) + \gamma_2 \cosh(\gamma_1 t)}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} \right] \\ B(t, T) &= \frac{1}{\sigma^2 \gamma_1} \frac{\kappa\theta\gamma_1 - \gamma_2\gamma_3 + \gamma_3[\sinh(\gamma_1 t) + \gamma_2 \cosh(\gamma_1 t)]}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} \\ C(t, T) &= -\frac{1}{2} \log[\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)] + \frac{1}{2}\kappa t + \frac{\kappa^2\theta^2\gamma_1^2 - \gamma_3^2}{2\sigma^2\gamma_1^3} \left( \frac{\sinh(\gamma_1 t)}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} - \gamma_1 t \right) \\ &\quad + \frac{(\kappa\theta\gamma_1 - \gamma_2\gamma_3)\gamma_3}{\sigma^2\gamma_1^3} \frac{\cosh(\gamma_1 t) - 1}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} \end{aligned}$$

where  $y = Y_0$  and  $\hat{s}_1 = \frac{1}{2}u^2\bar{\rho}^2 + \frac{1}{2}iu(1 - 2\kappa\rho/\sigma)$ ,  $\hat{s}_2 = iu\kappa\theta\rho\sigma^{-1}$ ,  $\hat{s}_3 = \frac{1}{2}iu\rho\sigma^{-1}$ .  $\phi_t(u)$  is regular in a neighborhood of the origin, so by Theorem 7.1.1 in Lukacs [Luk70],  $\phi_t(u)$  is also regular in the horizontal strip  $\{u \in \mathbb{C} : p_-(t) < u < p_+(t)\}$ , where

$$\begin{aligned} p_+(t) &= \sup_{p \geq 1} \mathbb{E}(e^{pX_t}) < \infty, \\ p_-(t) &= \inf_{p \leq 0} \mathbb{E}(e^{pX_t}) < \infty \end{aligned} \quad (3.3)$$

Note that  $p_{\pm}(t)$  is not the same as  $p_{\pm}$  as defined in the statement of the proposition, and we will show that  $p_{-}(t) \leq p_{-}$  and  $p_{+} \leq p_{+}(t)$  (see discussion above (3.6)).

Looking at the expressions for  $B, C, D$  on page 12 in [SZ99], we see that  $\phi_t(u)$  has a pole at  $u = -ip$  if and only if

$$\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t) = 0.$$

For  $p \in (p_{-}, p_{+})$  i.e. such that  $\Gamma(p) > 0$ , using that  $\gamma_1 = \Gamma(p)$  and  $-1/\gamma_2 = -\Gamma(p)/(\kappa - \rho p \sigma)$ , this equation is satisfied if  $t = t^*(p) = \frac{1}{\gamma_1} \tanh^{-1}(\frac{1}{\gamma_2}) = \frac{1}{\Gamma(p)} \tanh^{-1}(-\frac{\Gamma(p)}{\kappa - \rho p \sigma})$ . But negative  $t$ -values are physically meaningless, so our preliminary analysis would indicate that

$$T^*(p) = \begin{cases} \frac{1}{\gamma_1} \tanh^{-1}(\frac{1}{\gamma_2}) = \frac{1}{\Gamma(p)} \tanh^{-1}(-\frac{\Gamma(p)}{\kappa - \rho p \sigma}) & (\kappa - \rho p \sigma < 0) \\ +\infty & (\kappa - \rho p \sigma \geq 0) \end{cases}$$

where  $T^*(p) = \sup\{t : \mathbb{E}(e^{pX_t}) < \infty\}$  is the *moment explosion time*. We now first consider the case when  $p > 1$ . In this case, if  $\rho \leq 0$  then  $\kappa - \rho p \sigma > 0$  and for  $p \in (p_{-}, p_{+})$  we have that  $\Gamma(p) > 0$ , so  $T^*(p) = +\infty$ . Otherwise, if  $\rho > 0$ , then  $\kappa - \rho p \sigma < 0$  if  $p > p^*$  where  $p^* = \kappa/(\rho \sigma)$ . However

$$p^* - p_{+} = \frac{2\kappa - \rho\sigma - \rho\sqrt{4\kappa^2 - 4\kappa\rho\sigma + \sigma^2}}{2\rho\sigma\rho^2},$$

and using that

$$(2\kappa - \rho\sigma)^2 - \rho^2(4\kappa^2 - 4\kappa\rho\sigma + \sigma^2) = 4\kappa\rho^2(\kappa - \rho\sigma) > 0$$

we see that  $p^* > p_{+}$ , so it turns out that  $T^*(p) = \infty$  for *all*  $p \in (1, p_{+})$ . An almost identical calculation shows that  $T^*(p) = \infty$  for *all*  $p \in (p_{-}, 0)$ . Moreover, for  $p \in [0, 1]$ , from Jensen's inequality and the martingale property we have that  $\mathbb{E}(S_t^p) \leq S_0^p < \infty$  for all  $t$ . Thus we have shown that  $T^*(p) = \infty$  for all  $p \in (p_{-}, p_{+})$ , so the mgf of  $X_t$  is given by the analytic extension of  $\phi$  to the imaginary axis for  $p \in (p_{-}, p_{+})$ .

The expression for  $C(\cdot)$  in [SZ99] is given by

$$\begin{aligned} C(t, \hat{s}_1, \hat{s}_2, \hat{s}_3) &= -\frac{1}{2} \log[\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t) + \frac{1}{2}\kappa t] \\ &+ \frac{\kappa^2 \theta^2 \gamma_1^2 - \gamma_3^2}{2\sigma^2 \gamma_1^3} \left( \frac{\sinh(\gamma_1 t)}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} - \gamma_1 t \right) + \frac{(\kappa \theta \gamma_1 - \gamma_2 \gamma_3) \gamma_3}{\sigma^2 \gamma_1^3} \left( \frac{\cosh(\gamma_1 t) - 1}{\cosh(\gamma_1 t) + \gamma_2 \sinh(\gamma_1 t)} \right) \end{aligned}$$

where  $\gamma_1 = \sqrt{2\sigma^2 \hat{s}_1 + \kappa^2}$ ,  $\gamma_2 = (\kappa - 2\sigma^2 \hat{s}_3)/\gamma_1$  and  $\gamma_3 = \kappa^2 \theta - \hat{s}_2 \sigma^2$ , and for  $u = -ip$  and  $p \in (p_{-}, p_{+})$ , using that  $\gamma_1 = \Gamma(p) > 0$  and  $\cosh(\gamma_1 t) \sim \sinh(\gamma_1 t) \sim e^{\gamma_1 t}$  as  $t \rightarrow \infty$ , we obtain the following large-time behavior for  $C(t, \hat{s}_1, \hat{s}_2, \hat{s}_3)$ :

$$C(t, \hat{s}_1) \sim \frac{1}{2} \left[ \kappa + \frac{(p-1)p\theta^2\kappa^2}{\Gamma(p)^2} - \Gamma(p) \right] = t[V(p) + \frac{1}{2}p\rho\sigma] \quad (t \rightarrow \infty). \quad (3.4)$$

Letting  $t \rightarrow \infty$  and using that  $\coth(\gamma_1 t) \rightarrow 1$  as  $t \rightarrow \infty$ , we also find that

$$\begin{aligned} B(t, \hat{s}_1, \hat{s}_2, \hat{s}_3) &\sim -\frac{1}{\sigma^2 \gamma_1} [\gamma_3 - \kappa \theta \gamma_1] = O(1) \quad (t \rightarrow \infty), \\ D(t, \hat{s}_1, \hat{s}_2, \hat{s}_3) &\sim \frac{1}{\sigma^2} (\kappa - \gamma_1) = O(1) \quad (t \rightarrow \infty), \end{aligned}$$

and thus constitute higher order terms as  $t \rightarrow \infty$ , which we can ignore at the order we are interested in. Thus we have

$$\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{pX_t}) = V(p) \quad (3.5)$$

for  $p \in (p_{-}, p_{+})$ . This means that for  $p \in (p_{-}, p_{+})$  and  $t < \infty$  fixed, we have  $\mathbb{E}(e^{pX_t}) < \infty$ , so

$$\begin{aligned} p_{-}(t) &\leq p_{-}, \\ p_{+}(t) &\geq p_{+}. \end{aligned} \quad (3.6)$$

We now consider  $p \geq p_+$ . To this end we fix a  $q \in (1, p_+)$ ; then from the monotonicity of the  $L^p$  norm we have

$$(\mathbb{E}(e^{qX_t}))^{1/q} \leq (\mathbb{E}(e^{pX_t}))^{1/p}.$$

From this and (3.5) we obtain

$$V(q) = \liminf_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{qX_t}) \leq \liminf_{t \rightarrow \infty} \frac{1}{t} \log (\mathbb{E}(e^{pX_t}))^{q/p}.$$

But  $\forall K > 0$ , there exists a  $q(K) < p_+$  such that  $V(q) \geq K$ . Thus for  $t$  sufficiently large we have

$$(e^{Kt})^{p/q(K)} \leq \mathbb{E}(e^{pX_t})$$

or

$$K \leq K \frac{p}{q(K)} \leq \liminf_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{pX_t}).$$

Thus letting  $K \rightarrow \infty$  we see that  $\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{pX_t}) = +\infty$ . A similar analysis shows that  $\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(e^{pX_t}) = +\infty$  for  $p \leq p_-$ .

Differentiating  $V(p)$  we obtain

$$V'(p) = \frac{(2p-1)\theta^2\kappa^2 + 2(1-p)p\theta^2\kappa^2\Gamma'(p)}{\Gamma(p)^2} - \frac{1}{2}(\rho\sigma + \Gamma'(p))$$

and

$$\Gamma'(p) = \frac{\sigma(-2\kappa\rho + (1 + 2p(-1 + \rho^2))\sigma)}{2\Gamma(p)}.$$

Nothing that  $\Gamma(p_{\pm}) = 0$ , we see that  $V(p)$  and  $|V'(p)| \rightarrow +\infty$  as  $p \rightarrow p_{\pm}$  so  $V$  is essentially smooth, and  $V$  is lower semicontinuous.  $V_t(p) = \mathbb{E}(e^{pX_t})$  satisfies Assumption 2.3.2 in [DZ98] as  $t \rightarrow \infty$ , so by Lemma 2.3.9 in [DZ98]  $V$  is also convex, so from the Gärtner-Ellis Theorem (see Theorem 2.3.6 in [DZ98])  $X_t/t$  satisfies the LDP with good convex rate function  $I(x)$ .

We also have the upper bound

$$I(x) \leq p_+x \vee p_-x - V_{\min} < \infty$$

where  $V_{\min} = \inf_{p \in (p_-, p_+)} V(p) > -\infty$ . But a convex function is continuous on the interior of its domain, so  $I$  is continuous. Finally, from elementary calculations we find that the unique minimum of  $I$  occurs at  $x^* = (I')^{-1}(0) = V'(0)$ .  $\square$

#### 4. Call options and implied volatility

Let  $\mathbb{P}^*(A) = \frac{1}{S_0} \mathbb{E}(S_t 1_A)$  for  $A \in \mathcal{F}_t$  denote the *Share measure* ( $\mathbb{P}^*$  is a probability measure because  $S_t$  is a martingale by Proposition 3.1). From Girsanov's theorem, it is easily shown that

$$\begin{cases} d(-X_t) &= -\frac{1}{2}Y_t^2 dt - Y_t dW_t^{*1}, \\ dY_t &= [\kappa(\bar{\theta} - Y_t) + \rho\sigma Y_t] dt + \sigma dW_t^{*2} \\ &= \bar{\kappa}(\bar{\theta} - Y_t) dt + \sigma dW_t^{*2}, \end{cases} \quad (4.1)$$

where  $\bar{\kappa} = \kappa - \rho\sigma$ ,  $\bar{\theta} = \kappa\theta/(\kappa - \rho\sigma)$  and  $dW_t^{*1} dW_t^{*2} = \rho dt$  are independent  $\mathbb{P}^*$ -Brownian motions.

**Assumption 4.1.** *From here on we further assume that  $\bar{\kappa} = \kappa - \rho\sigma > 0$ , which ensures that  $Y_t$  is mean-reverting under  $\mathbb{P}^*$ .*

From (4.1), we have the following trivial corollary of Proposition 3.2.

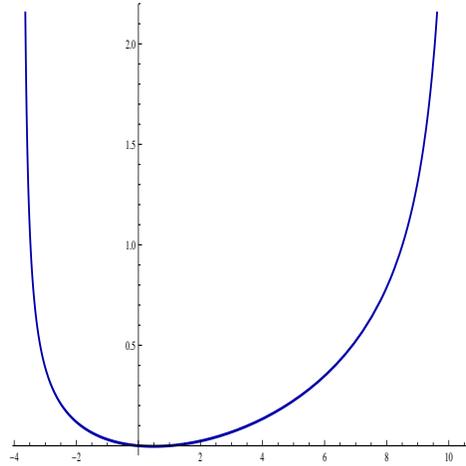


FIG 1. Here we have plotted  $V(p)$  for  $\kappa = 1.15, \theta = 0.1, \sigma = 0.2, \rho = -0.4$ .

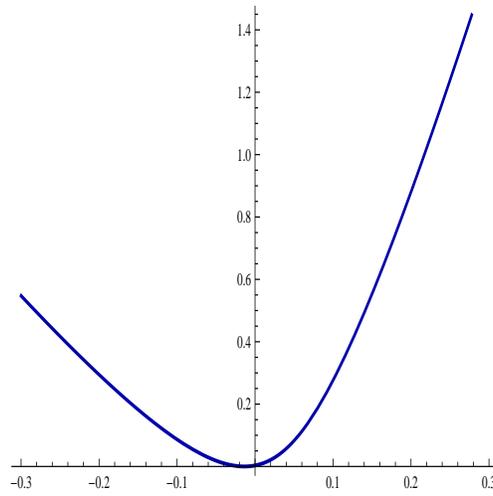


FIG 2. Here we have plotted the rate function  $I(x)$  for the same parameter values as above.

**Corollary 4.2.** For  $\kappa > \rho\sigma$ ,  $-X_t/t$  satisfies the LDP under  $\mathbb{P}^*$  as  $t \rightarrow \infty$  with a good convex continuous rate function  $I_S(x)$  given by the Fenchel-Legendre transform of

$$V_S(p) = V(p; \bar{\kappa}, \bar{\theta}, \sigma, -\rho)$$

and  $I_S$  is continuous and attains its minimum value uniquely at  $-x^+ = (V_S)'(0) = -\frac{1}{2}(\bar{\theta}^2 + \sigma^2/2\bar{\kappa})$ .

By Corollary 4.2 and the continuity of the rate function  $I_S$ , we obtain the following corollary, which will be used to characterize the large-time behaviour of call option prices.

**Corollary 4.3.**

$$\begin{aligned} \lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}^*(X_t > xt) &= -I_S(x) & (x > x_+), \\ \lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}^*(X_t < xt) &= -I_S(x) & (x < x_+). \end{aligned}$$

Recall that the payoff of a European call option with strike  $K$  is  $\mathbb{E}(S_t - K)^+$ , and the payoff of a European put option with strike  $K$  is  $\mathbb{E}(K - S_t)^+$ .

**Corollary 4.4.** We have the following large-time asymptotic behaviour for put/call options in the large-time, large log-moneyness regime:

$$\begin{aligned} -\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(S_t - S_0 e^{xt})^+ &= I_S(x) & (x \geq x_+), \\ -\lim_{t \rightarrow \infty} \frac{1}{t} \log [S_0 - \mathbb{E}(S_t - S_0 e^{xt})^+] &= I_S(x) & (x^* \leq x \leq x_+), \\ -\lim_{t \rightarrow \infty} \frac{1}{t} \log (\mathbb{E}(S_0 e^{xt} - S_t)^+) &= I_S(x) & (x \leq x^*) \end{aligned}$$

*Proof.* We first assume  $x > x_+$ , and recall that  $I_S(x)$  is non-decreasing for  $x > x_+$ . From Corollary 4.3, we know that for all  $\varepsilon > 0$  there exists a  $t^* = t^*(\varepsilon)$  such that for all  $t > t^*$  we have

$$\frac{1}{S_0} \mathbb{E}(S_t - S_0 e^{xt})^+ = \mathbb{P}^*(X_t > xt) - e^{xt} \mathbb{P}(X_t > xt) \leq \mathbb{P}^*(X_t > xt) \leq e^{-(I_S(x) - \varepsilon)/t}$$

which gives the upper bound for the call price. For the lower bound we have

$$\frac{1}{S_0} \mathbb{E}(S_t - S_0 e^{xt})^+ = \mathbb{E}^{\mathbb{P}^*}(1 - e^{xt} e^{-X_t})^+ = e^{xt} \mathbb{E}^{\mathbb{P}^*}(e^{-xt} - e^{-X_t})^+. \quad (4.2)$$

Observe that for any  $\delta > 0$ ,

$$\mathbb{E}^{\mathbb{P}^*}(e^{-xt} - e^{-X_t})^+ \geq \mathbb{E}^{\mathbb{P}^*}[(e^{-xt} - e^{-X_t})^+ \mathbf{1}_{\{-X_t < -xt - \delta\}}] \geq (e^{-xt} - e^{-xt - \delta}) \mathbb{P}^*(-X_t < -xt - \delta).$$

Combining this with (4.2) we have

$$\begin{aligned} \frac{1}{S_0} \mathbb{E}(S_t - S_0 e^{xt})^+ &\geq e^{xt} (e^{-xt} - e^{-xt - \delta}) \mathbb{P}^*(-X_t < -xt - \delta) \\ &= (1 - e^{-\delta}) \mathbb{P}^*(-X_t < -xt - \delta) \\ &= (1 - e^{-\delta}) \mathbb{P}^*(X_t/t > x + \delta/t) \\ &\geq (1 - e^{-\delta}) \mathbb{P}^*(X_t/t > x + \delta). \end{aligned}$$

Using Corollary 4.3 we get

$$\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(S_t - S_0 e^{xt})^+ \geq I_S(x + \delta).$$

This holds for all  $\delta > 0$ , so taking  $\lim_{\delta \rightarrow 0}$  and by the continuity of  $I_S(x)$  we obtain the first result that  $\lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E}(S_t - S_0 e^{xt})^+ = I_S(x)$ . The other cases follow similarly.  $\square$

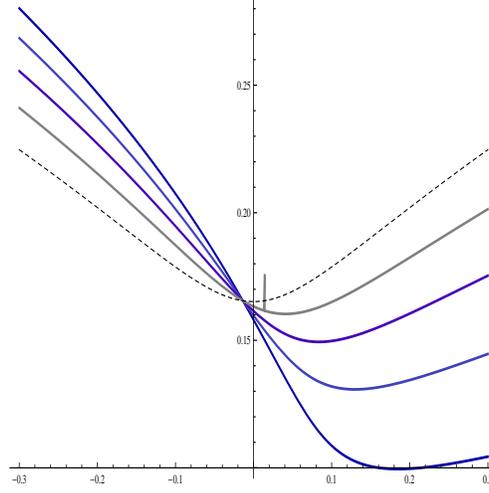


FIG 3. Here we have plotted the asymptotic implied volatility  $\hat{\sigma}(x)$  for  $\kappa = 1.15, \theta = 0.1, \sigma = 0.2$  and  $\rho = -.8, -.6, -.4, -.2$  and  $0$  (in blue, light blue, purple, grey and black dashed respectively).

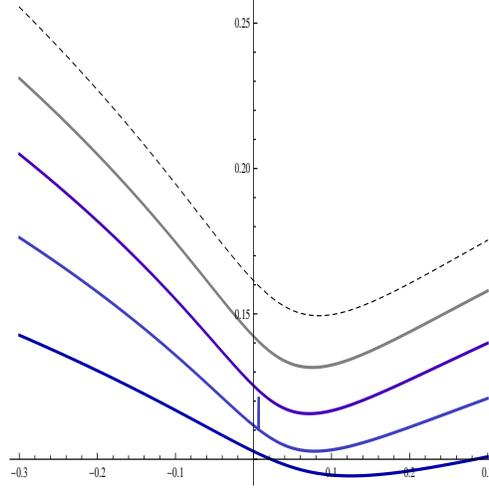


FIG 4. Here we have plotted the asymptotic implied volatility  $\hat{\sigma}(x)$  for  $\kappa = 1.15, \theta = 0.1, \rho = -0.4$  and  $\sigma = .04, .08, .12, .16$  and  $.2$  (in blue, light blue, purple, grey and black dashed respectively).

#### 4.1. Implied volatility

Using the same proofs as in Corollary 1.7 and Corollary 2.17 in [FJ11] for the Heston model (or Theorem 14 in [JKRM13] for a general affine model), we have the following asymptotic behaviour in the large-time, large log-moneyness regime, where  $\hat{\sigma}_t(x)$  is the implied volatility of a put/call option with strike  $S_0 e^{xt}$  for the correlated Stein-Stein model:

$$\hat{\sigma}_\infty(x)^2 = \lim_{t \rightarrow \infty} \hat{\sigma}_t^2(x) = \begin{cases} 2(2I(x) - x - 2\sqrt{I(x)^2 - I(x)x}) & (x \notin [x^*, x_+]) \\ 2(2I(x) - x + 2\sqrt{I(x)^2 - I(x)x}) & (x \in (x^*, x_+)) \end{cases}$$

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